Extended Methods and Results of the article "Hybridization could be a common phenomenon within the highly diverse lizard genus Liolaemus", by M. Olave, L.J. Avila, J.W. Sites Jr and M. Morando (2018), Journal of Evolutionary Biology.

## 1. Extended methods

### 1.1.Field work and laboratory procedures

Specimens were collected by hand or noose, sacrificed by a pericardiac injection of sodium tiopenthal Abbot $® /$ Pentovet $®$, dissected slightly to extract a sample of liver/muscle for molecular study, fixed in 10-20\% formalin, and later transferred to $70 \%$ ethanol. Tissues were stored in a freezer with $96 \%$ ethanol. Voucher specimens are deposited in the herpetological collections LJAMM-CNP of the Centro Nacional Patagónico, Puerto Madryn, Argentina
(CENPAT-CONICET, http://www.cenpat.edu.ar/nuevo/colecciones03.html), and the M. L. Bean Life Science Museum, Brigham Uoung University (BYU; http://mlbean.byu.edu/ResearchCollections/Collections/ReptilesandAmphibians.aspx).
For details of individuals used, see Appendix A. Genomic DNA was extracted using the Qiagen ${ }^{\circledR}$ DNeasy ${ }^{\circledR} 96$ Tissue Kit following the protocol provided by the manufacturer. For PCR and sequencing protocols we followed Morando et al. (2003) for the mitochondrial genes (cyt-b and 12S), Camargo et al. (2012) for the anonymous nuclear loci (ANL: A1D, A12D, A4B and A9C), and the nuclear protein-coding loci (NPCL): EXPH5, KIF24, MXRA5, (Portik et al. 2012), DNAH3, PRLR, PNN, SNCAIP (Townsend et al. 2008), and CMOS (Wiens et al. 1999). We amplified all nuclear genes using the touchdown PCR cycle described by Noonan and Yoder (2009), with standard reaction conditions (per sample: 2 $\mu \mathrm{l}$ dNTPs ( 1.25 mM ), $2 \mu \mathrm{l} 5 \mathrm{x}$ Taq buffer, $1 \mu \mathrm{l}$ each primer $(10 \mu \mathrm{M}), 1 \mu \mathrm{l} \mathrm{MgCl}(25 \mathrm{mM})$, and $0.1 \mu \mathrm{l}$ Taq DNA polymerase ( $5 \mathrm{U} / \mu \mathrm{l}$; Promega Corp., Madison, WI); 14 ml total reaction volume). All sequences were edited using the program Sequencher v4.8. ( ${ }^{\mathrm{TM}} \mathrm{Gene}$ Codes Corporation Inc. 2007), and aligned with MAFFT (Katoh et al. 2002) based on 100 tree rebuilding iterations and a maxirate of 100 . Protein-coding genes were translated to amino acids to check for codon errors. Missing data in all cases were coded as "?" ( $15.19 \%$ of the total dataset was missing). For each gene we selected the best-fitting model of evolution using JModelTest v2.1.1 (Darriba et al. 2012) using the Akaike Criterion Information (corrected) (AICc). In all nuclear genes, recombination was tested using RDP: Recombination Detection Program v3.44 (Martin and Rybicki 2000; Heath et al. 2006). Sequences are deposited in GenBank (Accession Nos. xxxx, to be completed upon ms acceptance), and alignments in Dryad data base (xxxx).

### 1.2. Gene trees and species trees

Gene tree inferences. We estimated each separated gene tree using the full matrix of 127 individuals and 14 loci using BEAST v1.6.1 (Heled and Drummond 2010). The two mitochondrial genes were concatenated to obtain a single gene tree, thus we obtained 13 independent gene trees in total. These trees were used for phylogenetic network estimation, as well as for presenting the mitochondrial gene tree in the main manuscript (Figure 2). Later we ran BEAST during $10 \times 10^{6}$ generations of MCMC and sampled at intervals of 1,000 generations with a $10 \%$ burnin. Convergence was diagnosed by observation of ESS values greater than 200 for each parameter estimated with Tracer v1.5 (Rambaut and Drummond 2009). The maximum credibility tree was obtained using TreeAnnotator.

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Species tree. The hybridization test used here requires a species tree topology as part of its model. Thus, we generated a species tree using only the focal species as part of the hybridization hypothesis. We included here the total of 14 loci and 95 individuals representing $L$. rothi, L. boulengeri, L. telsen and $L$. tehuelche. We ran the species tree analyses in *BEAST v1.6.1 (Heled and Drummond 2010) for $500 \times 10^{6}$ generations of MCMC, and sampled at intervals of 10,000 generations (burnin $10 \%$ ). Convergence was diagnosed by observation of ESS values greater than 200 for each parameter estimated. *BEAST also generates gene trees and saved them as outputs during the same analyses, thus we used these gene trees for XLC statistic calculations and hybridization test.

Divergence times estimation in the mitochondrial tree. We used substitution rates estimated for the mitochondrial regions for the genus Liolaemus: estimated as 0.006339 substitution per site per million years for 12 S , and 0.019355 substitution per site per million years for cyt-b. We used BEAST 1.6.1 and a strict clock prior independently for each mitochondrial region, estimated using a gamma distribution. For the 12 S the initial value was set $=0.006339$; shape $=1$; scale $=0.009$, which leads to a median equal to 0.006339 . For cyt-b the initial value was set $=0.019355$; shape $=1$; scale $=0.0174$, which leads to a median approximated equal to 0.019355 . We used the Yule process as the tree prior model, and ran a MCMC analysis for $100 \times 10^{6}$ generations, sampling every 10,000 intervals and a $10 \%$ burnin. Convergence was diagnosed by observation of ESS $>200$ for parameter estimation.

### 1.3.Testing the hybridization hypotheses

We set five $M$ parameters ( $M_{1}, M_{2}, M_{3}, M_{4}, M_{5}$ ) as is shown in Fig. 1 at the main manuscript (see also the species tree in newick format in Supplementary Material 2).
The simplest model has $M_{1,2,3,4,5}=0$ (strict coalescent model); and then we tested all parameters range between 1 to 5 . Thus, we tested models with different levels of gene flow between three different pairs of taxa: 1-L. rothi x L. tehuelche (current and past); 2-L. rothi x L. boulengeri + L. telsen (current and past); and 3-L. rothi and the common ancestor of (L. tehuelche + L. boulenger $+L$. telsen $)$. We calculate the likelihood of the data given each hypothesis, and selected the most likely model using likelihood ratio test. All used R functions are available in github.com/melisaolave/Olave_etal2017-MEE. We employed 13 gene trees (cyt-b and 12S mitochondrial loci were combined to infer a single gene tree) and a total of 93 samples of L. rothi, L. boulengeri, L. telsen and L. tehuelche.

Model construction. Models included a species tree topology and branch lengths (inferred as described above in section 1.2 in Supplementary Material) and migration parameters ( $M$ $=\mathrm{N}_{\mathrm{e}} m$; where $\mathrm{N}_{\mathrm{e}}$ is the effective population size and $m$ is the proportion of migrants).

Branch lengths were converted to coalescent units to make them compatible with the ms program ( $m s$ branch length units $=4 \mathrm{~N}_{\mathrm{e}}$ generations), which was implemented by 1
multiplying the branch lengths by the population parameter $\bar{\theta}\left(\theta=4 \mathrm{~N}_{\mathrm{e}} \mu ; \mathrm{N}_{\mathrm{e}}\right.$ is the effective population size and $\mu$ the mutation rate) for simulating nuclear gene trees, and for mitochondrial gene tree by $\frac{\mathbf{4}}{\boldsymbol{\theta}}$. This was done, by setting brlength.correction in get.topo2ms function.

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We used Lamarc v2.1.8 (Kuhner 2006) to estimate $\theta$ values in a Bayesian inference of 31,000 MCMC steps, with a $10 \%$ burnin. Diagnosis of convergence was made by observing ESS values equal or greater than 200. Here, $\theta$ was estimated $=0.05399$.
Simulations, extra lineage approximation and the likelihood ratio test were performed using the R functions available in github.com/melisaolave/Olave_etal2017-MEE, described by (Olave et al. 2017a). Specifically, we used the function get.topo2ms() to load the species tree and obtain input commands to ms program (Hudson 2002) program, and the ms() function to simulate 10,000 gene trees per locus to infer the distribution of expected extra lineages given each proposed model. The number of extra lineages were counted using the sptree.vs.genetree() function to automatize the process using Phylonet package (Than and Nakhleh 2009). We obtained the likelihood of the data given each model using the get.Likelihood() function. The model was then selected taking the parameter estimations and comparing the likelihood of nested models using a likelihood ratio test implemented in the function Lratiotest().

### 1.4.Morphological comparisions

Morphologies were quantified using 10 morphometric variables standardly used in taxonomic studies of the genus, including: snout-vent length (SLV, distance from the tip of the snout and the posterior margin of the precloacal scales), axilla-groin distance (AGD, distance from the armpit of the right front leg to the anterior insertion of the hindlimb), hand length (HaL, distance from the base of the wrist and base of the nail of the third digit; measured ventrally), radius-ulna length (RUL, distance from the hand to the elbow), humerus length (HL, distance from the elbow and the shoulder; ventrally measured), foot length ( $\mathbf{F o L}$, distance from the base of the heel to the base of the nail of the fourth digit; measured ventrally), tibio-fibula length (TFL, distance from knee to the internal angle with the foot), femur length ( $\mathbf{F L}$, distance from the knee to its pelvic insertion), and knee-knee distance (KKD, distance between knees bent a right angles to the abdomen, measured ventrally).

We also included the following 15 color pattern variables: vertebral line melanism (VLM, shape of the dorsal spots), number of spots at the left vertebral line (NSLVL), number of spots at the right vertebral line (NSRVL), presence of a white spot below SVL spots (PWSVL), paraventral melanism (PVM), number of left paraventral spots (NLPS), number of right paraventral spots (NRPS), presence of a white spot below PVM spots (PWSPL), mean line (ML, presence of line between MLV and MPV), scapular melanism (SM, spots at both sides (left and right) between the neck and the shoulders), gular melanism (GM, ventrally observed), ventral melanism (VM), ventral color (VC, color ventrally observed), forelimb melanism (FM), and hindlimb melanism (HM).

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## 2. Extended results:

### 2.1. Supplementary tables

| Locus | Substitution model <br> selected |
| :---: | :---: |
| cyt-b | HKY+I+G |
| 12 S | GTR $+\mathrm{I}+\mathrm{G}$ |
| CMOS | HKY |
| DNAH3 | HKY+I |
| EXPH5 | HKY+G |
| KIF24 | HKY+I |
| A12D | GTR+G |
| A1D | HKY+G |
| A4B | HKY+G |
| A9C | HKY+I |
| MXRA5 | HKY+I |
| PNN | HKY+I |
| PRLR | HKY+I |
| SNCAIP | HKY+I |
| Table S1: Substitution models selected by jModeltest. |  |


| Summary statistic | Locus |  |
| :---: | :---: | :---: |
|  | Cyt-b | 12 S |
| Median | 0.025715 | 0.01626 |
| HPD | $0.0030635-0.0061598$ | $0.0011281-0.025398$ |
| Mean | 0.028901 | 0.01922 |
| Standard deviation | 0.00027147 | 0.00010991 |

Table S2: Results of posterior estimation of substitution rates for the boulengeri-rothi complexes.

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| Species/group | Centroids <br> Axis 1 | Axis 2 |
| :--- | :--- | :--- |
| i.mt (i) | -0.81 | -1.88 |
| i.mt (ii) | -0.84 | -0.69 |
| L. rothi | -1.18 | 0.61 |
| L. telsen | 1.61 | -0.67 |
| L. tehuelche | 1.61 | 0.75 |
| L. boulengeri | 1.77 | -0.27 |

Table S3: Centroid coordinates estimated by the multivariate discriminant analysis (Fig. S1).

| Group | i.mt (i) | i.mt (ii) | L. rothi | L. telsen | L. tehuelche | L. boulengeri | Total | Error (\%) |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| i.mt (i) | 12 | 0 | 1 | 0 | 0 | 1 | 14 | 14.29 |
| i.mt (ii) | 2 | 2 | 2 | 0 | 0 | 0 | 6 | 66.67 |
| L. rothi | 2 | 3 | 34 | 2 | 1 | 1 | 43 | 20.93 |
| L. telsen | 0 | 0 | 0 | 5 | 0 | 1 | 6 | 16.67 |
| L. tehuelche | 0 | 0 | 0 | 0 | 14 | 3 | 17 | 17.65 |
| L. boulengeri | 0 | 0 | 0 | 1 | 1 | 15 | 17 | 11.76 |
| Total | 16 | 5 | 37 | 8 | 16 | 21 | 103 | 20.39 |

Table S4: Cross-classification table from the discriminant analysis (Fig. S1). The rows represent the groups provided a priori and the columns represent the assignments based on the discriminant analysis. The proportion of miss-classifications is shown in the Error (\%) column. While most of the individuals were classified in their original groups, all groups showed some proportion of classification error. For example, some of the $L$. rothi samples were classified as part of the boulengeri species complex, with one L. rothi sample each associated with L. tehuelche, L. boulengeri and L. telsen, while the reciprocal cases were not inferred. Individuals of the i.mt (i) and (ii) groups are miss-classified in both directions, as well as with $L$. rothi group.

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### 2.2. Supplementary figure



Fig. S1: Discriminate analysis plot estimated for the quantitative variables. Centroids are shown in Table S3 and cross-classification inferences are presenting in Table S4. The i.mt (i) and (ii) represent the hybrid candidates.

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